Third trimester placental volume and biometry measurement: A method-development study

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ABSTRACT

Objectives: To test the hypothesis that third trimester placental biometry and volume can be measured by two-dimensional (2D) and three-dimensional (3D) ultrasound in utero, determining which method of measurement was most strongly correlated with true placental size ex vivo.

Methods: Singleton pregnancies underwent placental ultrasound within seven days of delivery (n = 87, 29-35-41-5 weeks). Length and width (linear and curvilinear) and depth were estimated. Placental volume (PV) was estimated using 2D ellipse and shell techniques and 3D rotational (15° and 30° rotation angles) and multiplanar (5 and 10 mm slicing intervals) techniques. Measurements were compared to their true correlates following delivery. Intra- and inter-observer reliabilities of candidate placental size estimates were assessed by intraclass correlation coefficient (ICC).

Results: Curvilinear placental length ($R_s = 0.24, p = 0.031$), width ($R_s = 0.31, p = 0.0056$) correlated well with ex vivo measurements. All methods of PV estimation were related to ex vivo volume ($R_s > 0.32, p < 0.01$) but not placental weight ($p > 0.05$); 30° rotational estimation demonstrated the strongest biological correlation ($R_s = 0.40, p = 0.0004$). Intra- and inter-observer placental size measurements intraclass correlation coefficients were suboptimal (0.59–0.70 and 0.10–0.58 respectively).

Discussion: We have demonstrated that it is possible to obtain information about the size of the third trimester placenta in utero using 2D and 3D ultrasound. However it is essential that the reliability (particularly interobserver reliability) of these estimates is improved prior to prospective studies to determine their predictive value.

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1. Introduction

Smaller ex vivo placental size, with or without increase in fetoplacental ratio (FPR), is observed in pregnancies ending in stillbirth [1,2], fetal growth restriction (FGR) [3,4] and reduced fetal movement pregnancies with adverse outcome [5] compared to normal outcome pregnancies. Sonographically detectable placental growth restriction precedes FGR by several weeks [6], thus assessment of in utero placental size, alone or in relation to fetal size, may improve prediction of adverse pregnancy outcome.

Two-dimensional (2D) ultrasound measures of placental diameter and thickness have been used as indicators of high-risk pregnancies [7–11]. First trimester sonographic placental volume (PV) using the three-dimensional (3D) ultrasound technique Virtual Organ Computer Aided analysis (VOCAL) is smaller in pregnancies ending in delivery of small for gestational age (SGA) infants [12,13] early-onset FGR and hypertensive disorders [14–16]. In contrast late-onset FGR and preeclampsia pregnancies failed to show an appreciable significant difference in first trimester VOCAL PV [16–18]. However, no studies have examined the accuracy of these techniques, or whether there is any meaningful correlation between in utero sonographic placental biometry measurements and direct measurements of ex vivo placental size after delivery.

In contrast, a 2D PV technique has been correlated to placental weight (PW) [19]. It is not known whether PW and PV are consistently related; therefore this correlation may prove inappropriate. In the second trimester this technique demonstrates reasonable...
specificity (91%) but low sensitivity (19%) for SGA birth [20]; pregnancies with PV < 25th centile are twice as likely to experience adverse outcome [21].

This study investigated the relationship between placental volume and weight, and tested the hypothesis that placental biometry and volume can be accurately and reproducibly measured in utero in third trimester pregnancies. Here we define accuracy as the ability of in vivo measurement to relate to ex vivo measurement and reproducibility to describe the variability of the measurement. Biological relevance (relationship of the measure to an outcome of interest) is not tested in this study.

2. Materials and methods

Women with non-anomalous singleton pregnancies of ≥28 weeks gestation undergoing third trimester ultrasound examination (for assessment of reduced fetal movements, suspected or confirmed FGR or confirmation of fetal presentation) gave written informed consent to participate in the ethically approved study (11/NW/0650, Greater Manchester North West Research Ethics Committee) by undergoing additional ultrasound measurements and placental donation.

2.1. Sonographic assessment of placental size and shape

Ultrasound examinations were conducted by a single sonographer (LH) using a Voluson E6 with a RA4B 4–8 Hz curvilinear probe (GE Healthcare). Fetal weight was estimated (EFW) according to the Hadlock C formula [22]. Using 2D ultrasound the placental was located and its longest plane identified. The ultrasound probe was angulated to include as much placenta as possible and to minimise acoustic shadowing from the fetus. A 2D image and 3D volume (85° sweep) of the placenta were captured in this plane. The probe was rotated 90°, the longest perpendicular plane identified and a further 2D image was captured. This procedure was repeated three times. 2D images were analysed in real time. 3D images were analysed offline using 4Dview version 5.0 (GE Healthcare) ultrasound image analysis software. Throughout the manuscript, the prefix ‘est’ refers to an in utero estimate of a particular aspect of placental size. Accuracy refers to the existence of a significant statistical correlation between estimated and true measurements. Biological relevance (relationship of the measure to an outcome of interest) is not tested in this study.

2.2. Modelling of placental shape and tissue density

Placentas were trimmed of their extra-placental membranes and umbilical cord. The PW and PV (measured by volume displacement [23]) were recorded and the placentas were photographed, chorionic plate facing upward alongside a scale bar. Placental depth (D) was measured directly at the apparent deepest point of the placenta. Using Image ProPlus version 6.0 (Media Cybernetics UK, Marlow, UK) placental photographs were analysed to quantify placental length (L; longest diameter of the placenta), width (W; longest diameter perpendicular to the placental length) and average diameter (A). These measurements were incorporated into formulae for the volume of an ellipse \( \frac{4}{3} \pi A^2 \times 0.5 \) and a circular cylinder \( \pi \times (0.5 \times 0.5 \times 0.5) \). The modelled ex vivo PV determined by each formula was then correlated to true ex vivo PV. The model that best approximated true PV was carried forward into subsequent analyses. Tissue density was expressed as the ratio of PW to PV.

2.3. Correlation of sonographic and true placental biometry, volume, weight and fetoplacental ratio

Subgroup analysis of those placentas delivered within seven days of ultrasound examination was performed to test the accuracy of sonographic placental measurements. Systematic and random errors were calculated for the most accurate method of estimating each placental size measure (as determined by the highest correlation coefficient of statistically significant measurement methods).

2.3.1. Placental biometry

\( \text{est} \) from images of the longest plane of the placenta and \( \text{est} \) (from images of the longest perpendicular plane of the placenta) were estimated in three ways; (i) a straight line (or two straight lines meeting at the angle of the placenta if the placenta was particularly curved) through the placenta from tip to tip (Fig. 1A); (ii) a curvilinear line along the maternoplacental interface (Fig. 1B) [25], and (iii) a curvilinear line through the middle of the placenta (Fig. 1C). For 2D measurements, if the placental length or width could not be fully captured in any single image, real-time extrapolation based on the data ascertained from movement of the probe along the placenta was permitted. \( \text{est} \) was estimated at the visibly deepest point of the placenta, perpendicular to its plane (Fig. 1D). The most accurate method of estimating \( \text{est} \) and \( \text{est} \), were carried forward into further analyses.

2.3.2. Placental volume and weight

\( \text{est} \) (from images of the longest plane of the placenta) and \( \text{est} \) (from images of the longest perpendicular plane of the placenta) were estimated in three ways; (i) a straight line (or two straight lines meeting at the angle of the placenta if the placenta was particularly curved) through the placenta from tip to tip (Fig. 1A); (ii) a curvilinear line along the maternoplacental interface (Fig. 1B) [25], and (iii) a curvilinear line through the middle of the placenta (Fig. 1C). For 2D measurements, if the placental length or width could not be fully captured in any single image, real-time extrapolation based on the data ascertained from movement of the probe along the placenta was permitted. \( \text{est} \) was estimated at the visibly deepest point of the placenta, perpendicular to its plane (Fig. 1D). The most accurate method of estimating \( \text{est} \) and \( \text{est} \), were carried forward into further analyses.

2.3.3. Fetoplacental ratio

FPR was assessed to examine placental efficiency, dividing EFW by \( \text{est} \) and \( \text{est} \) to respectively generate fetoplacental volume and weight ratios (\( \text{est} / \text{est} \) and \( \text{est} / \text{est} \)).

2.4. Assessment of sonographic reliability

Throughout the study data was collected regarding reproducibility, by analysis of intra- and inter-observer reliability of placental sonographic measures in a series of scans conducted in triplicate (\( N = 46 \)) by two sonographers (LH and LS) following the same methodologies described above. Three measurement sets were obtained (\( \text{LH} \), \( \text{LS} \), \( \text{LH} \)) from each participant, with each sonographer blinded to the values obtained in each previous assessment. Intra-observer reliability was assessed by comparison of \( \text{LH} \) and \( \text{LH} \) values, whilst inter-observer reliability was assessed by
comparison of LH\textsubscript{1} and LS values [31].

2.5. Statistical analysis

The study cohort was divided into subgroups by scan to delivery interval using a cut off of seven days. Subgroups were compared to assess for selection bias; data were compared by Mann-Whitney \textit{U} Test (for continuous data) and Chi squared test (for categorical data). Sonographic accuracy was assessed in women giving birth within 7 days by Spearman Rank Correlation. The strength of that association was determined by the highest Rs value of each set of measurements; where equal, the most accurate method of estimation was determined to be the one with smallest systematic error (Systematic error: (predicted \textendash true)/true. Random error: standard deviation of systematic errors). The most accurate method of estimating each measurement was taken forward into subsequent analyses. Reliability (intra- and inter-observer) was assessed as the coefficient of variance (CoV; standard deviation of measures/average of measures), intra-class correlation coefficient (ICC (3,1)), bias and 95% limits of agreement. An ICC >0.75 was considered good [32]. Statistical analysis was carried out using Prism 6 for Mac OS X (Graphpad Software Inc., San Diego, USA) and statistical significance was determined by a \textit{p} value < 0.05.

Based on an expected correlation of 0.68 between \textit{estPV} and \textit{PV} in term pregnancies and an up to 37.5\% failure to obtain third trimester measurements [19] a minimum of 22 placentas would be required to detect a significant correlation of similar magnitude with power of 80\% at the level of \textit{p} < 0.05. Given the local proportion of pregnancies delivering within seven days of assessment for reduced fetal movements (32.6, unpublished data) and a 10\% loss to follow up rate, we estimated that at least 120 participants would need to be recruited to obtain sufficient matched in vivo and

Fig. 1. Measurement of placental biometry using two-dimensional ultrasound. Length and width were measured by three methods; straight line (A), curvilinear line at maternal-placental interface (B) and curvilinear line through the placenta (C). Placental depth was measured perpendicular to the placenta at the maximal point (D). Broken line represents line of measurement.

Fig. 2. Measurement of placental volume using two- and three-dimensional ultrasound. Schematic representations of placental volume estimation by two-dimensional (A&B) and three-dimensional (C&D) ultrasound with accompanying axial and sagittal orthographic planes. Ellipse (A) and shell (B) volumes were calculated according to formulae for the volume of an ellipse and elliptical shell respectively. Rotational (C) and multiplanar (D) volumes were calculated using in-built formulae within 4D view v.5 (GE Healthcare) ultrasound image analysis software after tracing the placental outline at pre-specified rotation or slicing intervals. Broken line represents line of measurement.
3. Results

3.1. Recruitment and assessment of biological correlation

Placentas from 129 participants were received; these pregnancies formed the whole study cohort. Table 1 summarises their maternal and pregnancy characteristics. Eighty-seven (67.4%) pregnancies delivered within seven days of ultrasound examination; in vivo and ex vivo measurements were compared in this group. One or more 2D placental measurements were not obtained in two cases (2.3%) and 3D volumes were deemed inadequate due to excessive movement artefact or inadequate visualisation in four cases (4.6%). EFW was highly significantly correlated with true birth weight (R = 0.77, p < 0.0001) indicating that the sonographic accuracy in general was good.

3.2. Ex vivo placental volume modelling and tissue density

Each shape approximation modelled PV with highly significant relationship to the true PV (R = 0.74, p < 0.0001 for each). However, systematic bias differed significantly; ellipse (6.4% underestimation), elliptical cylinder (42.2% overestimation) and circular cylinder (182.8% overestimation). PW and PV were highly significantly correlated (R = 0.96, p < 0.0001 for each). How-ever, systematic bias differed significantly; ellipse (6.4% underestimation), elliptical cylinder (6.4% underestimation) and circular cylinder (182.8% overestimation).

3.3. Correlation of sonographically estimated placental biometry, volume, weight and fetoplacental ratio

3.3.1. Placental biometry

Table 2 summarises the performance of each described method of measuring 2D placental biometry. Statistically significant relationships were observed between estL and L using only method 3, whilst for estW both methods 1 and 3 each correlated with W with statistical significance; once more method 3 performed best. estD was significantly related to the D. Systematic and random errors for each measurement were significant, particularly for estD (Systematic error: estL = −14.76%, estW = −13.35%, estD = 97.83%. Random error: estL = 17.14%, estW = 18.32%, estD = 49.66% respectively).

3.3.2. Placental volume and weight

Table 2 also summarises the performance of each method of estimating PV, estPV by the 2D ellipsoidal model (using estL and estW by method 3 above) was more accurate than the previously published shell method. All 3D estPV methods demonstrated higher

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**Table 1**
The placental ultrasound study cohort.

<table>
<thead>
<tr>
<th>Scan to delivery interval (days)</th>
<th>Whole cohort</th>
<th>Subgroup-analysis</th>
<th>p</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Any</td>
<td>&lt;7</td>
<td>&gt;7</td>
</tr>
<tr>
<td>N</td>
<td>129</td>
<td>87</td>
<td>42</td>
</tr>
</tbody>
</table>

**Maternal Characteristics**

| Age (years)                  | 29.2 (25.0–33.1) | 30.8 (26.0–33.5) | 29.0 (23.6–33.2) | 0.31 |
| Ethnicty:                     |               |                  |                  | 0.34 |
| Caucasian                     | 93 (72.1%)     | 63 (72.4%)       | 30 (71.4%)       |     |
| Asian                         | 17 (13.2%)     | 11 (12.6%)       | 6 (14.2%)        |     |
| Black                         | 9 (7.0%)       | 8 (9.2%)         | 1 (2.4%)         |     |
| Other                         | 10 (7.8%)      | 5 (5.7%)         | 5 (11.9%)        |     |
| BMI (kg/m²)                   | 26.0 (23.1–30.2) | 25.5 (23.1–29.5) | 28.1 (22.9–31.6) | 0.30 |
| Parity (number)               | 1 (0–1)       | 0 (0–1)          | 0 (0–1)          | 0.12 |

**Pregnancy Characteristics**

| Gestation at Scan (weeks+days) | 38−0 (37.2–40.1) | 39−0 (38.2–40.3) | 36−0 (32.3–38.2) | <0.0001 |
| Placental site:                |               |                  |                  |       |
| Anterior                       | 39 (30.2%)     | 28 (32.2%)       | 11 (26.1%)       |       |
| Lateral                        | 45 (34.9%)     | 40 (46.0%)       | 5 (11.9%)        |       |
| Posterior                      | 15 (11.6%)     | 9 (10.3%)        | 6 (14.3%)        |       |
| Fundal                         | 12 (9.3%)      | 7 (8.0%)         | 5 (11.9%)        |       |
| Not defined                    | 18 (14.0%)     | 3 (3.4%)         | 15 (35.7%)       |       |
| Scan to delivery interval (days)| 3 (1–7)       | 2 (0–4)          | 20 (12–54)       | <0.0001 |
| Individualised birth weight centile | 36 (14.5–61.2) | 36 (14.3–63.2) | 48 (15.5–60) | 0.56 |
method of \( \text{estPW} \) or \( \text{estPV} \) demonstrated statistical correlation with PW, although several approached statistical significance. This may be a result of the relatively small sample size studied and suboptimal reproducibility (see below), however we believe this currently prohibits use of PW to validate \( \text{estPV} \), and prevents generation of a dimensionless third trimester placental quotient (FPRw) [34].

While estimated and true placental size values are correlated, they are not equal with \( \text{estL}, \text{estW} \), and 3D \( \text{estPV} \) being smaller and \( \text{estD} \) and 2D \( \text{estPV} \) larger than their ex vivo correlates. \( \text{estD} \) and \( \text{estPV} \) may be inflated by placental blood [35] whilst the other measures may be reduced by “missed” placental tissue. Furthermore the strength of the relationship is lower here than predicted from previous studies [19]. This implies that sonographic placental measurements should be compared against in vivo, rather than ex vivo reference curves.

Our study further suggests that development of clinically useful in vivo reference curves may be impaired by the intra- and inter-observer variability in these measures. The effect of the operator “learning curve” on reliability was not examined, however performance may improve with experience, particularly in relation to extrapolation if contemporaneous feedback was provided regarding accuracy of placental measurement. 2D ultrasound reliability data are limited; a singular study of second trimester placental biopsy reports very high reliability (ICCs > 0.92) [36]. At such early gestation the majority of placentas should be fully visualised in a single image, removing the need for real time extrapolation. Extrapolation is felt likely to contribute to the lower reliability estimates for the 2D placental estimates in this third trimester study. However the specific contribution of extrapolation variability cannot be ascertainment in the current study.

A significant contribution of “missed tissue” to the suboptimal reliability of third trimester measures may also be inferred as, using the same technique, (and without extrapolation) other researchers have reported much higher intra- and inter-observer ICCs > 0.88 using VOCAL \( \text{estPV} \) in the first and second trimesters [29,30,37], and between 12 and 40 weeks gestation (majority of measurements obtained < 28 weeks gestation) [26] than those demonstrated in this third trimester study. However, other researchers employing robust methodologies have also demonstrated similar reliability (ICC 0.59 with wide limits of agreement) using this technique even in the first trimester [38].

It is not yet known whether differences in the size of term pregnancy placentas that exceed full visualisation in a single image are clinically relevant. Indeed depending on factors such as placental site, maternal habitus, fetal size and gestation the impact of excluding “missed” placental tissue in the assessment of placental size might have a disproportionate effect on the usefulness of a test of placental size if real-time extrapolation was not performed. Whether the reliability demonstrated in our study (ICC < 0.75) is sufficient to detect relatively subtle differences (e.g. <3 cm difference in placental length and width [5]) in placental size remains to be seen, although both Pomorski et al. [39] and Artunc Ulkumen et al. [28] were able to detect a 92 cm³ reduction in VOCAL \( \text{estPV} \) between third trimester FGR and control pregnancies. This should be the subject of future work in this area, alongside comparative assessment of the clinical utility of these measures against other assessments of in utero placental structure and function including umbilical and uterine artery Doppler impedance and placental hormonal assessment.

The strength of this study is the like-for-like correlation of placental size estimation techniques to their true biological correlates with a short scan-to-delivery interval. This resulted in development of placental measurements that may be tested prospectively to subsequently determine relationship to pregnancy.

### 3.4. Assessment of sonographic reliability

\( \text{Table 3} \) summarises the reliability indices of each placental measurement. For all measures the variability in measurements (intra- and inter-) was suboptimal with no ICC > 0.75. Inter-observer variability exceeded intra-observer variability with Bland Altman plots demonstrating wide limits of agreement for all measures (Fig. 3). The degree of bias was unaffected by placental size (\( p > 0.05 \)).

<table>
<thead>
<tr>
<th>Measure</th>
<th>CoV (%)</th>
<th>ICC</th>
<th>0070</th>
</tr>
</thead>
<tbody>
<tr>
<td>Length</td>
<td>Intra</td>
<td>8.3</td>
<td>0.68</td>
</tr>
<tr>
<td></td>
<td>Inter</td>
<td>14.0</td>
<td>0.096</td>
</tr>
<tr>
<td>Width</td>
<td>Intra</td>
<td>9.5</td>
<td>0.70</td>
</tr>
<tr>
<td></td>
<td>Inter</td>
<td>10.8</td>
<td>0.37</td>
</tr>
<tr>
<td>Depth</td>
<td>Intra</td>
<td>12.4</td>
<td>0.59</td>
</tr>
<tr>
<td></td>
<td>Inter</td>
<td>13.0</td>
<td>0.58</td>
</tr>
<tr>
<td>Volume</td>
<td>Intra</td>
<td>14.4</td>
<td>0.66</td>
</tr>
<tr>
<td></td>
<td>Inter</td>
<td>26.2</td>
<td>0.54</td>
</tr>
</tbody>
</table>

Intra- and inter-observer reliability was assessed in a series scans (N = 46) by two observers. Variability between observations is displayed as the coefficient of variance (CoV) and intraclass correlation coefficient (ICC).
outcome in the future. The study assessed placentas in advanced gestation, irrespective of placental site, maternal body mass index and fetal size, in an ethnically diverse cohort, making the study findings generally applicable in a wide range of health care settings. There is little reason to suspect that these methodologies would not remain valid at earlier gestations.

The primary limitations of the study are that each method of placental size estimation demonstrated relatively poor intra- and inter-observer reliability of placental size estimate replicates as demonstrated by wide scatter in a series (N = 46) of repeated measures and between measures by two observers; length (A), width (B), depth (C), volume (D). For intra-observer reliability, the average of both readings from one observer is shown on the x-axis and the between reading difference (expressed as a percentage of the average of both readings) on the y-axis. For inter-observer reliability the average of both observer's readings is shown on the x-axis and the between observer difference (expressed as a percentage of the average of both readings) on the y-axis. The bias between observations is depicted by a solid grey line (where not visible, overlying y = 0); the 95% limits of agreement are depicted by a broken line.

Fig. 3. Intra-observer and Inter-observer reliability of placental size estimates. Bland-Altman plots demonstrate no systematic bias (p > 0.05), but suboptimal within observer reliability of placental size estimate replicates as demonstrated by wide scatter in a series (N = 46) of repeated measures and between measures by two observers; length (A), width (B), depth (C), volume (D). For intra-observer reliability, the average of both readings from one observer is shown on the x-axis and the between reading difference (expressed as a percentage of the average of both readings) on the y-axis. For inter-observer reliability the average of both observer's readings is shown on the x-axis and the between observer difference (expressed as a percentage of the average of both readings) on the y-axis. The bias between observations is depicted by a solid grey line (where not visible, overlying y = 0); the 95% limits of agreement are depicted by a broken line.
inter-observer reliability, particularly in ESL measurement. This implies that significant refinement of the technique is required prior to clinical application. Further investigation of the contribution of extrapolation to both reliability, statistical correlation and subsequently to pregnancy outcome prediction is essential before this aspect of the technique is accepted or rejected. Other potential limitations of the study include the cross-sectional nature, sample size and high-risk population, which prevent generation of centile charts.

5. Conclusions

With increasing interest in antenatal placental assessment to identify potentially compromised pregnancies, this study provides evidence that placental size can be estimated in the third trimester with statistical correlation to its true size. The biological relevance of these results remains untested, but may be limited by suboptimal reliability of these measurements. Thus, reliability of in utero placental size estimates needs to be improved before prospective studies to determine if placental size assessment is useful in prediction of pregnancy outcome.

Statement of author contributions

The project was conceived by EDJ, AEPH, CPS and LH, and methodologies planned by LH with expert supervision from EDJ (sonography). LS performed ultrasound scans for analysis of inter-observer reliability. LH performed all other ultrasound scans and all other analyses. Statistical analysis and manuscript preparation was performed by LH. All authors were involved in writing the paper and had final approval of the submitted and published versions.

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